

Let us return to our model. We have a solution of Kolmogorov's forward equation

$$\frac{\partial f_N(t, \xi_1, \dots, \xi_N)}{\partial t} = N^2 \mathcal{A}_N f_N$$

where  $f_N$  is the density with respect to the invariant measure given by  $\exp[-\sum \phi(\xi_i)]$ . We saw that the Dirichlet form satisfies.

$$\int \frac{\sum |(D_i - D_{i+1})f_N|^2}{f_N} e^{-\sum \phi(\xi_i)} d\xi \leq \frac{C}{N}$$

If we denote the marginal in a block of length  $\ell$  around  $x$  by  $f_{N,x\ell}(\xi_1, \dots, \xi_\ell)$  and by  $\bar{f}_{N<\ell}$  the average

$$\bar{f}_{N,\ell} = \frac{1}{N} \sum_x f_{N,x,\ell}$$

then the Dirichlet form

$$\int \frac{\sum_{1 \leq i \leq \ell-1} |(D_{i+1} - D_i)\bar{f}_{N,\ell}|^2}{\bar{f}_{N,\ell}} e^{-\sum \phi(\xi_i)} d\xi \leq \frac{C\ell}{N^2}$$

With a log Sobolev constant of  $\ell^2$ , we can apply log Sobolev inequality on each hyperplane  $\sum_i x_i = \ell a$ . If we write  $\bar{f}_{N,\ell} e^{-\sum \phi(\xi_i)} d\xi = \bar{G}_{N,\ell}(da) \bar{\nu}_{N,\ell}(a, d\xi)$  and similarly  $e^{-\sum_i \phi(\xi_i)} = \phi_{N,\ell}(a) \bar{\mu}_{N,\ell}(a, d\xi)$

$$\int H(\bar{\nu}_{N,\ell}(a, \cdot) | \bar{\mu}_{N,\ell}(a, \cdot)) \bar{G}_{N,\ell}(da) \leq c \frac{\ell^3}{N^2}$$

If we can prove a conditional version of Cramér's Large deviation result, i.e. with  $\lambda = h'(a)$ , and

$$\bar{F}(a) = \frac{1}{M(\lambda)} \int e^{\lambda \xi - \phi(\xi)} F(\xi) d\xi$$

$$P\left[\left|\frac{1}{n} \sum_i F(\xi_i) - \bar{F}(a)\right| \geq \delta \mid \frac{1}{n} \sum \xi_i = a\right] \leq \exp[-nC(\delta)]$$

where  $C(\delta) > 0$ , then it is not hard to see that with  $\ell = N\epsilon$ ,

$$\int_{\left|\frac{1}{n} \sum_i F(\xi_i) - \bar{F}(a)\right| \geq \delta} \bar{f}_{N,\ell}(\xi) d\xi \leq c \frac{(N\epsilon)^3}{N^3 \epsilon} = c\epsilon^2$$

This will allow replacing averages like

$$\frac{1}{N} \sum J\left(\frac{x}{N}\right) F(\xi)$$

with

$$\frac{1}{N} \sum J\left(\frac{x}{N}\right) \bar{F}\left(\frac{1}{2N\epsilon} \sum_{y:|y-x| \leq N\epsilon} \xi_y\right)$$

with a small  $\epsilon$ . The conditional version of Cramér's theorem is not hard to prove.

Let us start with i.i.d.r.v's with density  $g(\xi)$ . Let  $g_n(\xi)$  be the density of  $\frac{1}{n} \sum \xi_i$ . Then

$$g_n(a) = \exp[-nh(a) + o(n)]$$

is the density version of the Large deviation result. Let us take this for granted. We want to compute

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log E[e^{\rho \sum_i F(\xi_i)} | \sum_i \xi_i = na] = \psi(\rho, a)$$

Suppose we define

$$g_\rho(\xi) = \frac{1}{T(\rho)} e^{\rho F(\xi)} g(\xi)$$

normalized to be a probability distribution. Then the distribution of  $\frac{1}{n} \sum \xi$  under the new distribution is given by a density  $g_{n,\rho}(a)$  that will satisfy

$$g_{n,\rho}(a) = \exp[-nh(\rho, a) + o(n)]$$

It is not hard to see that

$$\psi(\rho, a) = \log T(\rho) - h(\rho, a) + h(a)$$

and

$$h(\rho, a) = \sup_{\theta} [\theta a - \log M(\rho, \theta)]$$

with

$$M(\rho, \theta) = \frac{1}{T(\theta)} \int e^{\rho F(\xi) + \theta \xi} g(\xi) d\xi$$

Getting exponential error estimate is just the differentiability of  $\psi(\rho, a)$  at  $\rho = 0$  which is obvious. Moreover

$$\psi_\rho(0, a) = \frac{T'(0)}{T(0)} - h_\rho(0, a)$$

and

$$h_\rho(0, a) = -\frac{M_\rho(0, \theta)}{M(0, \theta)} = \frac{T'(0)}{T(0)} - \frac{1}{M(0, \theta)} \int e^{\theta \xi} F(\xi) g(\xi) d\xi$$

with  $\theta = h'(a)$ . Finally

$$\psi_\rho(0, a) = \frac{1}{M(0, \theta)} \int e^{\theta \xi} F(\xi) g(\xi) d\xi$$

with  $\theta = h'(a)$ . Finally let us prove the density version of Cramér's theorem.

We write with  $g_\theta(\xi) = \frac{1}{M(\theta)} e^{\theta \xi} g(\xi)$  and  $\theta = h'(\xi)$ , so that

$$\int \xi g_\theta(\xi) d\xi = a$$

$$\begin{aligned} \int_{\bar{\xi}=a} g(\xi_1) \cdots g(\xi_n) d\xi &= M(\theta)^n e^{-na} \int_{\bar{\xi}=a} g_\theta(\xi_1) \cdots g_\theta(\xi_n) d\xi \\ &= M(\theta)^n e^{-na} g_{n,\theta}(a) \end{aligned}$$

Since now  $a$  is the mean, the density version of CLT yields

$$g_{n,\theta}(a) \simeq \frac{1}{\sqrt{2\pi n C(a)}}$$

Density version of CLT. It is sufficient to prove with mean 0 and variance 1

$$\lim_{n \rightarrow \infty} \frac{1}{2\pi} \int \widehat{g}\left(\frac{t}{\sqrt{n}}\right)^n e^{-itx} dt \rightarrow \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

It is not hard to get a bound of the form

$$|\widehat{g}(t)| \leq (1 + ct^2)^{-c}$$

and for  $n \gg 1$ ,

$$\left(1 + c \frac{t^2}{n}\right)^{-nc} \leq (1 + c't^2)^{-1}$$